

Let's generalize these "results"

Heuristic Derivation of EOF or Principal Component Analysis

Assume we have M time series $u_i(t) = u_i(t, x_i)$

We define an EOF as a set of functions that allow the following transformation

$$u_i(t, x_i) = \sum_{n=1}^M A_n(t) \cdot \phi_n(x_i)$$

vector space
of data

vector space of
statistical "modes"

where

n denotes the n-th EOF mode

$A_n(t)$ denotes the amplitude of the n-th EOF mode (eigenfunction)

$\phi_n(x_i)$ is the pattern (eigenvector) of the n-th EOF mode

i denotes the variable i-th variable

The functions $A_n(t)$ and $\phi_n(x_i)$ are subject to the following orthogonality conditions

$$(1.) \quad \sum_{i=1}^M \phi_n(x_i) \phi_m(x_i) = \delta_{nm} = \begin{cases} 0 & n \neq m \\ 1 & n = m \end{cases}$$

$$(2.) \quad \frac{1}{K} \sum_{k=1}^K A_n(t_k) A_m(t_k) = \delta_{nm} \frac{1}{K} \sum_{k=1}^K A_n^2(t_k)$$

$$\langle A_n(t) A_m(t) \rangle = \delta_{nm} \lambda_n$$

I will show next that the above definitions and constraints constitute an eigenvalue problem for the cross-correlation matrix.

Consider time series at two locations, i.e., $u_i(t, x_i)$ and $u_j(t, x_j)$. The product of these data are

$$u_i(t, x_i) \cdot u_j(t, x_j) = \sum_{n=1}^M \sum_{m=1}^M A_n(t) \phi_n(x_i) A_m(t) \phi_m(x_j)$$

taking the time (or ensemble) average, i.e., $\langle \cdot \rangle \equiv \frac{1}{K} \sum_{k=1}^K \cdot(t_k)$

$$\langle u_i(t, x_i) \cdot u_j(t, x_j) \rangle = \sum_{n=1}^M \sum_{m=1}^M \langle A_n(t) A_m(t) \rangle \phi_n(x_i) \phi_m(x_j)$$

using (2.)

$$= \sum_{n=1}^M \sum_{m=1}^M \delta_{nm} \lambda_n \phi_n(x_i) \phi_m(x_j)$$

$$= \sum_{n=1}^M \lambda_n \phi_n(x_i) \phi_n(x_j)$$

For an arbitrary mode m at location j we have

$$\sum \phi_m(x_j) \langle u_i u_j \rangle = \phi_m(x_j) \sum_{n=1}^M \lambda_n \phi_n(x_i) \phi_n(x_j) \quad m=1, 2, \dots, M$$

$$= \sum_{n=1}^M \lambda_n \phi_n(x_i) \phi_n(x_j) \phi_m(x_j)$$

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And summing over all locations j

$$\sum_{j=1}^M \phi_m(x_j) \langle u_i, u_j \rangle = \sum_{j=1}^M \sum_{n=1}^M \lambda_n \phi_n(x_i) \phi_n(x_j) \phi_m(x_j)$$

$m=1, 2, \dots, M$

$$= \sum_{n=1}^M \lambda_n \phi_n(x_i) \sum_{j=1}^M \phi_n(x_j) \phi_m(x_j)$$

$$\text{using (1.)} \quad = \sum_{n=1}^M \lambda_n \phi_n(x_i) \delta_{nm}$$

$$\sum_{j=1}^M \phi_m(x_j) \langle u_i, u_j \rangle = \lambda_m \phi_m(x_i) \quad m=1, 2, \dots M$$

Which defines an eigenvalue problem

It is the #

$$\begin{matrix} i & j \\ \left. \begin{array}{c} m=1 = i \\ m=2 = i \\ \vdots \\ \vdots \end{array} \right| & \left. \begin{array}{c} \langle u_1, u_1 \rangle \quad \langle u_1, u_2 \rangle \dots \langle u_1, u_M \rangle \\ \langle u_2, u_1 \rangle \quad \langle u_2, u_2 \rangle \dots \langle u_2, u_M \rangle \\ \vdots \quad \vdots \quad \ddots \quad \vdots \\ \langle u_M, u_1 \rangle \quad \langle u_M, u_2 \rangle \dots \langle u_M, u_M \rangle \end{array} \right| \end{matrix} \left| \begin{array}{c} \phi_m(x_1) \\ \phi_m(x_2) \\ \vdots \\ \phi_m(x_M) \end{array} \right| = \hat{\mu}_m$$

of time series

$$= \begin{pmatrix} \lambda_m & 0 & \dots & 0 \\ 0 & \lambda_m & & \\ \vdots & \vdots & \ddots & \\ 0 & 0 & \dots & \lambda_m \end{pmatrix} \begin{pmatrix} \phi_m(x_1) \\ \phi_m(x_2) \\ \vdots \\ \phi_m(x_M) \end{pmatrix} = \lambda_m \begin{pmatrix} \phi_m(x_1) \\ \phi_m(x_2) \\ \vdots \\ \phi_m(x_M) \end{pmatrix}$$

The set of functions $\{\phi_1, \phi_2, \dots, \phi_M\}$ form the basis vectors of a "new" vector space such that the data can be represented as

$$u_i(x_i, t_2) = \sum_{n=1}^M A_n(t_2) \phi_n(x_i)$$

where

$$A_n(t_2) = \sum_{i=1}^M u_i(x_i, t_2) \phi_n(x_i)$$

We still need to find the eigenfunctions ~~or~~ or eigenvectors $\phi_n(x_i)$

~~or~~ → singular value decomposition (better use TRED2 + TQL4 of Numerical Recipes)
Let us first, however, have a close look at the matrix

$$R_{ij} = \langle u_i u_j \rangle = \langle u_i(t, x_i) u_j(t, x_j) \rangle = \frac{1}{K} \sum_{k=1}^K u(t_k, x_i) u(t_k, x_j)$$

The diagonal elements R_{ii} represent the variance of variable i :

$$\begin{aligned} R_{ii} &= \langle u_i u_i \rangle = \left\langle \sum_{n=1}^M A_n \phi_n(x_i) \cdot \sum_{n=1}^M A_n \phi_n(x_i) \right\rangle \\ &= \left\langle \sum_{n=1}^M \sum_{m=1}^M A_n \phi_n(x_i) A_m \phi_m(x_i) \right\rangle \\ &= \sum_{n=1}^M \sum_{m=1}^M \langle A_n A_m \rangle \phi_n(x_i) \phi_m(x_i) \\ &= \sum_{n=1}^M \sum_{m=1}^M \delta_{nm} \lambda_n \phi_n(x_i) \phi_m(x_i) \end{aligned}$$

$$R_{ii} = \sum_{n=1}^M \lambda_n \phi_n^2(x_i)$$

summing up the diagonal elements gives the total variance

$$\sum_{i=1}^M R_{ii} = \sum_{i=1}^M \sum_{n=1}^M \lambda_n \phi_n^2(x_i)$$

$$\text{using (1)} \quad = \sum_{n=1}^M \lambda_n \sum_{i=1}^M \phi_n^2(x_i)$$

$$\sum_{i=1}^n \phi_n(x_i) \phi_n(x_i) = \delta_{nn}$$

$$= \sum_{n=1}^M \lambda_n$$

Hence the trace of the matrix R_{ii}

$$T = \sum_{i=1}^M R_{ii} = \sum_{i=1}^M \lambda_i = \sum_{i=1}^M \langle A_n^2 \rangle$$

↑
using (2)

thus

and we can interpret the eigenvalues λ_n as the variances associated with the n -th mode because the new "variables"

orthogonal

$A_n(t)$ can be viewed as amplitudes whose ^{sum of} square equals the trace of the covariance matrix R_{ij} which is the sum of the variances of the individual records $u_i(t)$, i.e.,

$\frac{\lambda_n}{\sum_{n=1}^M \lambda_n}$ is the % variance explained by the n -th mode